



A Non-stationary Service Curve Model for Performance Analysis of Transient Phases in Cellular Networks

Nico Becker

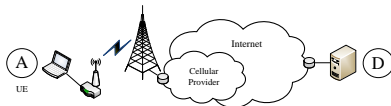
Markus Fidler



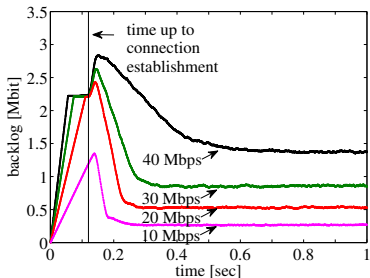
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INSTITUT FÜR KOMMUNIKATIONSTECHNIK



Testbed LTE



Backlog LTE

- ▶ Transient phase due to DRX mode
- ▶ In DRX mobile devices enter sleep phases to save energy
- ▶ Waking up causes additional delays
- ▶ Trade-off between energy saving and additional delay
- ▶ Relevant for safety-critical applications



Stationary vs. Non-stationary Service Curves

Deterministic Sleep Scheduler

Random Sleep Scheduler

Measurement-Based Estimation

Rate Scanning

Burst Response

Minimal Probing

Measurements in LTE



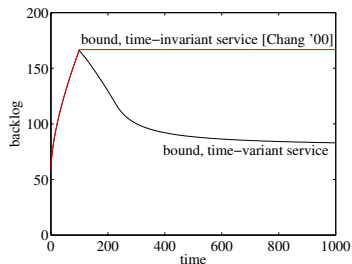
Consider an service process $S(t)$. The process is stationary, if

$$P[S(\tau, t) \leq x] = P[S(\tau + \delta, t + \delta) \leq x], \quad (1)$$

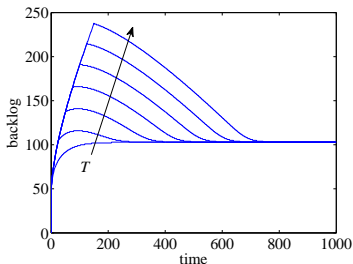
for any $\tau, t, \delta \geq 0$, i.e. the probability to see a certain amount of service in an interval does not depend on the time instance at which the interval starts but only on the duration of the interval.



- ▶ Consider a transmitter and a receiver that if idle go to a sleep state according a defined protocol.
- ▶ Wake up is scheduled deterministically, T time units after entering sleep state.
- ▶ The transmission rate in sleep state is zero and otherwise it is R .



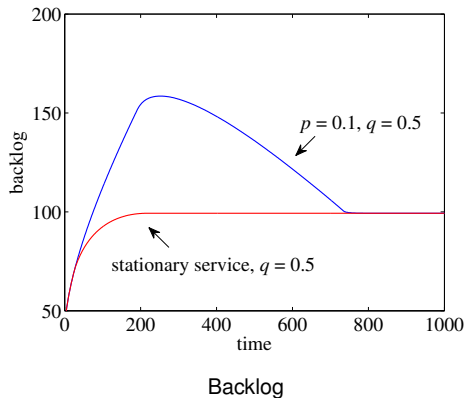
time-variant vs. time-invariant



$$T = \{0, 25, \dots, 150\}$$



- ▶ Consider a transmitter and a receiver that if idle go to a sleep state according a defined protocol.
- ▶ Wake up is scheduled randomly T time units after entering sleep state, i.e., T is geomatrically distributed with parameter p .
- ▶ The transmission rate in sleep state is zero and otherwise a Bernoulli increment process with parameter q .



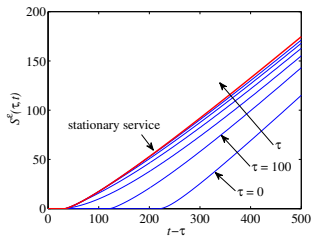


Let $S(\tau, t)$ be a bivariate random service process. Then, any function $S^\varepsilon(\tau, t)$ that satisfies

$$P[S(\tau, t) \geq S^\varepsilon(\tau, t), \forall \tau \in [0, t]] \geq 1 - \varepsilon, \quad (2)$$

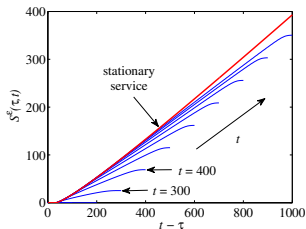
is a **non-stationary service curve**

for all $t \geq 0$, where $\varepsilon \in (0, 1]$ is the underflow probability.



$$\tau = \{0, 100, \dots\}$$

(fixing τ)



$$t = \{0, 100, \dots\}$$

(fixing t)



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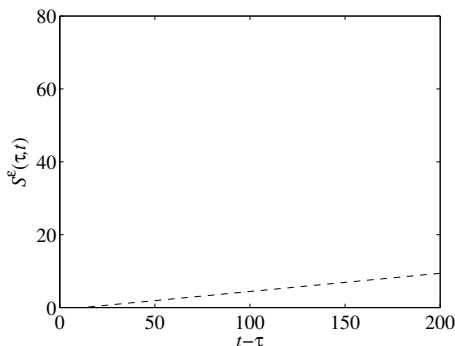
Measurements in LTE



- ▶ Uses constant rate probes $A(t) = rt$, for a set of probes $r \in \mathbb{R}$
- ▶ $S(\tau, t) \geq \max_{r \in \mathbb{R}} \{r(t - \tau) - B(r, t)\}$
- ▶ Repeat measurements and take backlog quantile $B^\xi(r, t)$
- ▶ $S^\varepsilon(\tau, t) = \max_{r \in \mathbb{R}} \{r(t - \tau) - B^\xi(r, t)\}$
- ▶ $\varepsilon = \sum_{r \in \mathbb{R}} \xi$ (Union bound)

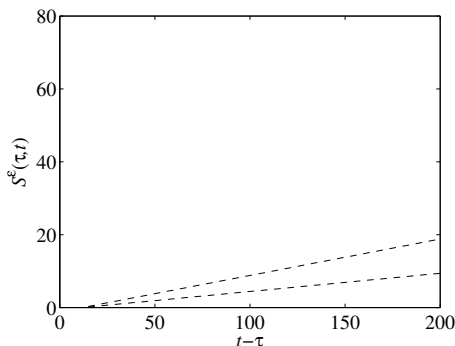


- ▶ Example for the random sleep scheduler, with $p = 0.1$ and $q = 0.5$.
- ▶ For every rate $r \in \{0.05, 0.1, \dots, 0.5\}$ we get 10^5 backlog samples.
- ▶ $\xi = 10^{-4}$ so that $\epsilon = \sum_{r \in \mathcal{R}} \xi = 10^{-3}$



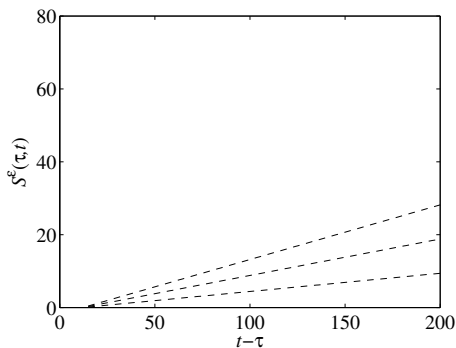


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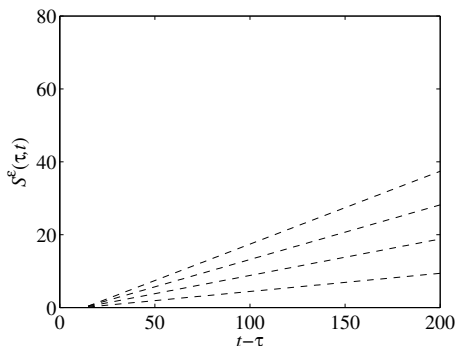


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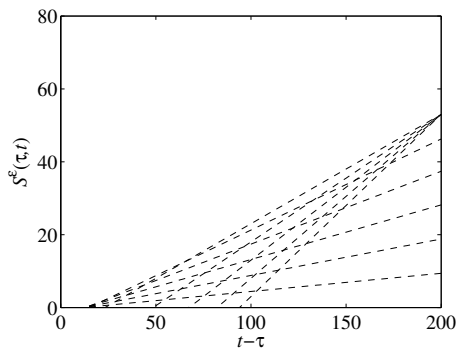


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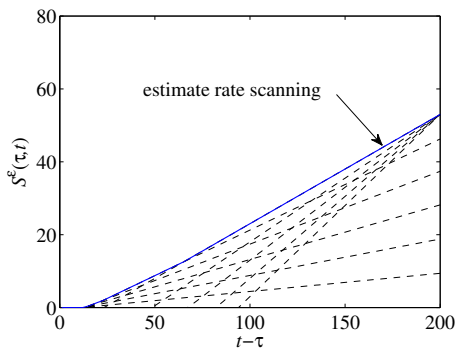


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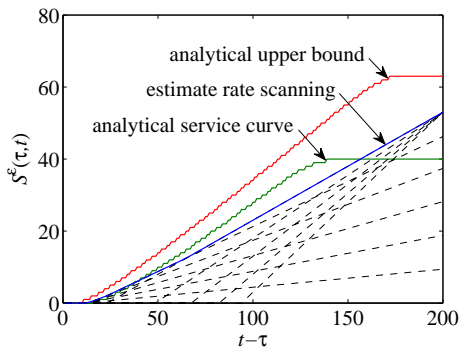


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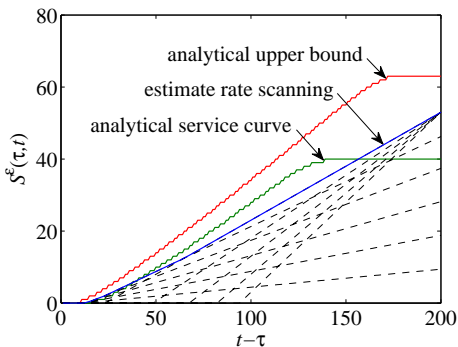


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The service curve cannot recover the non-convex part of the analytical results.





- ▶ Uses canonical probes for system identification, i.e.,

$$A(\tau) = \delta(\tau) = \begin{cases} 0 & \text{for } \tau = 0, \\ \infty & \text{for } \tau > 0. \end{cases}$$



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- ▶ $D(t) = \inf_{\tau \in [0, t]} \{\delta(\tau) + S(\tau, t)\} = S(0, t)$
- ▶ For additive service processes: $S(\tau, t) = S(0, t) - S(0, \tau)$

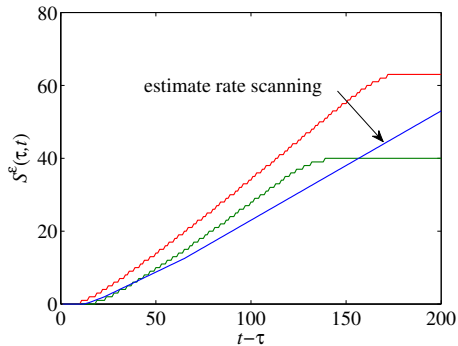


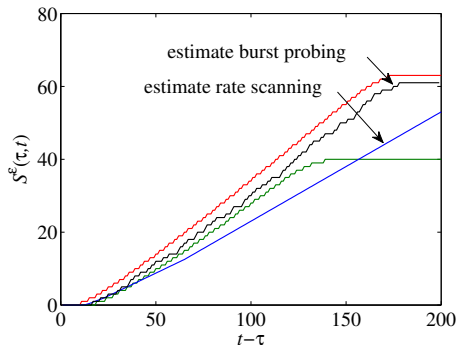
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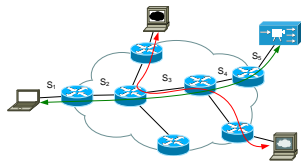
- ▶ $D(t) = \inf_{\tau \in [0, t]} \{\delta(\tau) + S(\tau, t)\} = S(0, t)$
- ▶ For additive service processes: $S(\tau, t) = S(0, t) - S(0, \tau)$
- ▶ Repeat measurements to get the set of all feasible sample Ω
- ▶ Remove the worst-cases and compute from the remaining set Ψ the non-stationary service curve

$$S_{br}^{\varepsilon}(\tau, t) = \inf_{\psi \in \Psi} \{S_{\psi}(\tau, t)\} \quad (3)$$





- ▶ Burst probes cause non-linear behavior of certain systems.
- ▶ Preempt other traffic, resulting in a too optimistic service estimate.



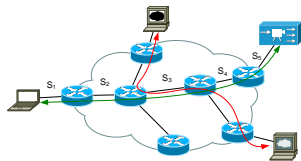
For additive and univariate
service S^i ($i = 1, 2, \dots, n$):

$$S^{net}(\tau, t) = S^1 \otimes S^2 \otimes \dots \otimes S^n(\tau, t)$$

Lemma (Super-additivity of \otimes)

Given two bivariate functions $f(s, t)$ and $g(s, t)$ for $t \geq s \geq 0$ where $f(t, t), g(t, t) = 0$ for all $t \geq 0$. Define $h(s, t) = f \otimes g(s, t)$.

If f and g are super-additive, then h is super-additive.



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If f and g are super-additive, then h is super-additive.

$$\Rightarrow S^{net}(\tau, t) \leq S^{net}(0, t) - S^{net}(0, \tau)$$

Note, that it includes additive processes, as well!



- ▶ We seek to find the minimal probe that satisfies for a fixed t
$$D(t) = \inf_{\tau \in [0, t]} \{A_{mp}(\tau) + S(\tau, t)\} = S(0, t),$$
- ▶ i.e., the minimal probe that allows estimating the service from observations of the departures.
- ▶ The minimal probe is $A_{mp}(\tau) = S(0, t) - S(\tau, t)$.
- ▶ For any other larger or smaller probe it leads to a lower service.
- ▶ We do not know $S(\tau, t)$ in advance



1. Compute the burst response estimate , i.e., $S_{br}^{\varepsilon}(\tau, t)$
2. Use $S_{br}^{\varepsilon}(\tau, t)$ to compute the minimal probe, i.e.,

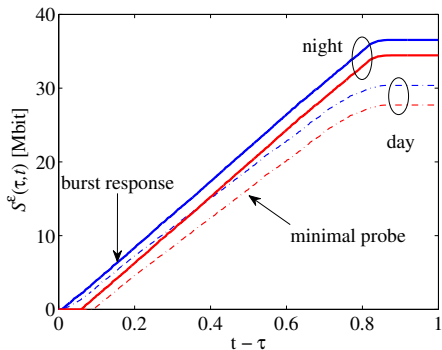
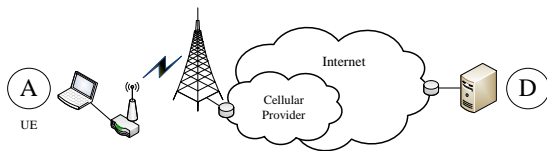
$$\tilde{A}_{mp}(\tau) = S_{br}^{\varepsilon}(0, t) - S_{br}^{\varepsilon}(\tau, t) \quad (4)$$

and repeat the measurements to get the service for the minimal probe, $S_{mp}^{\varepsilon}(\tau, t)$.



For $\tilde{A}_{mp}(\tau) = S_{br}^{\varepsilon}(0, t) - S_{br}^{\varepsilon}(\tau, t)$ we conclude that $B^{\varepsilon}(t)$ observed by minimal probing is a measure of accuracy that separates the conservative estimate of minimal probing from the possibly too optimistic estimate of burst probing, i.e.,

$$S_{mp}^{\varepsilon}(\tau, t) = S_{br}^{\varepsilon}(\tau, t) - B^{\varepsilon}(t) \quad (5)$$





- ▶ Analysis of non-stationary service curves
- ▶ Evaluated the effect on transient phases (also in comparison to stationary service curves)
- ▶ Devised a novel two-phase method to obtain an accurate service curve estimate
- ▶ Simulation results confirmed the fidelity of the approach
- ▶ Measurements in LTE show that the method is applicable in practice.



[BF '15] Becker, Fidler. : A Non-stationary Service Curve Model for Performance Analysis of Transient Phases.

in *Proc. of ITC 27*, Sep 2015

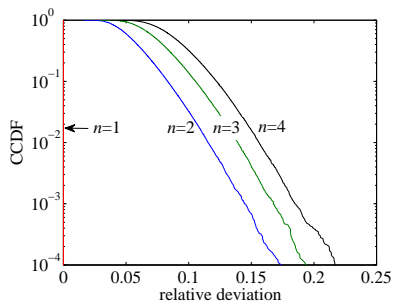
[Chang '00] C.-S. Chang. : Performance Guarantees in Communication Networks. *Springer-Verlag*, 2000

[LFL '14] Lübben, Fidler, Liebeherr. : Stochastic Bandwidth Estimation in Networks With Random Service.

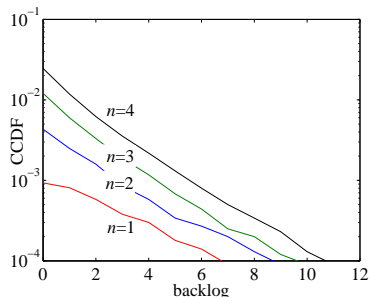
IEEE/ACM Trans. Netw., vol. 22, no.2, pp. 484-497, Apr 2014

[LFV '10] Liebeherr, Fidler, Valaee. : A System Theoretic Approach to Bandwidth Estimation.

IEEE/ACM Trans. Netw., vol. 18, no.4, pp. 1040-1053, Aug 2010



(a) Deviation from additivity



(b) Accuracy of minimal probing

Figure: Network of n systems with random sleep scheduling in series. (a) The network service process deviates from additivity. (b) Minimal probing achieves small backlogs, corresponding to a high accuracy of the estimate.



Consider an service process $S(t)$. The process is stationary, if

$$P[S(\tau, t) \leq x] = P[S(\tau + \delta, t + \delta) \leq x], \quad (6)$$

for any $\tau, t, \delta \geq 0$, i.e. the probability to see a certain amount of service in an interval does not depend on the time instance at which the interval starts but only on the duration of the interval.



Let $S(\tau, t)$ be a bivariate random service process. Then,

- i. any function $S^\varepsilon(t)$ that satisfies

$$P[S(\tau, t) \geq S^\varepsilon(t - \tau), \forall \tau \in [0, t]] \geq 1 - \varepsilon, \quad (7)$$

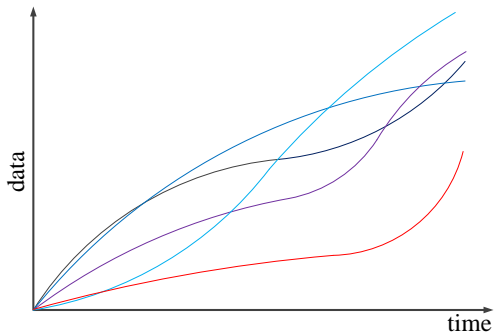
is an **ε -effective service curve**

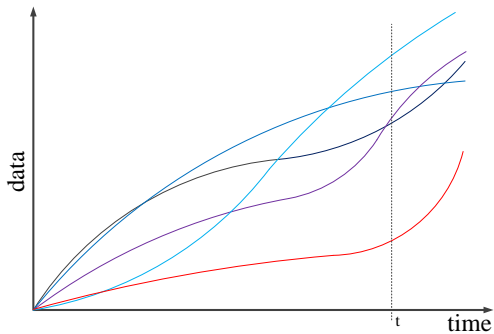
- ii. any function $S^\varepsilon(\tau, t)$ that satisfies

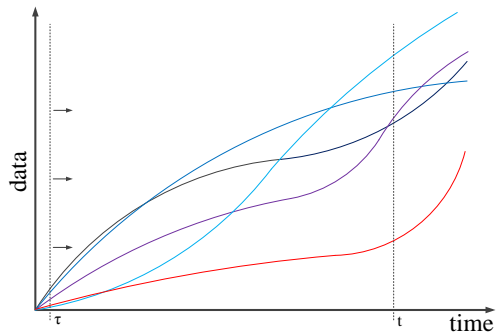
$$P[S(\tau, t) \geq S^\varepsilon(\tau, t), \forall \tau \in [0, t]] \geq 1 - \varepsilon, \quad (8)$$

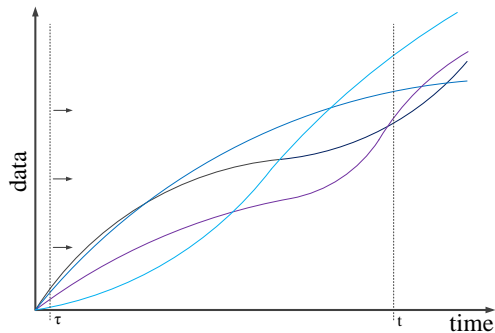
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